

**ALTERNATIVE APPROACHES TO MEASURING
THE COST OF EDUCATION**

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Introduction

Over 20 years after the *Serrano v. Priest* (1971) decision by the California Supreme Court sparked intense debate over school finance equity, the topic remains at the forefront of the education reform debate in many states. Over the past two decades, a number of states have faced law suits over the equity of their school finance system and several states have been forced to make changes. In the last several years, a new round of court cases has challenged traditional equity standards and solutions implemented in response to past court challenges. This paper addresses a central issue in this debate, namely educational cost differences across school districts, that has been virtually ignored by the courts and left out of recent reform efforts.

For the most part, the school finance debate has focused on differences in school district fiscal capacity, and aid formulas typically make some effort to compensate low-capacity districts. Much less attention has been paid to the other side of the school district budget, where cost differences have a major impact on educational outcomes. The courts have focused on the equalization of expenditure per pupil and not on adjusting expenditure to achieve equal educational outcomes. Cost adjustments made by states tend to involve ad hoc adjustment factors, such as "weighted pupil measures" to account for high-cost students and scale factors to compensate small, rural school districts.¹ Aid formulas based on these cost factors are likely to under-adjust for cost differences, and indeed may even magnify existing disparities instead of easing them.

Over the last decade, several scholars have developed methods for constructing educational cost indices. In this literature, the need to account for education cost factors is widely acknowledged, but scholars disagree about the best way to define and measure costs. As we use the term, "cost" refers to the expenditure or outlay needed by a district to provide a specified level of education attainment or outcome, not to actual expenditure. In other words, *cost refers to the value of the resources a district must consume in the production of a given level of student achievement*. Cost differentials reflect both the costs of inputs and the harshness of the production environment.² Actual expenditure, on the other hand, reflects the influence not only of cost factors, but also of demand factors, such as tax price, and of institutional factors.

Our objectives in this paper are to develop a method for estimating a comprehensive district-level educational cost index that builds on the existing literature and can be implemented with available data and then to estimate this index using data for New York State. Although we do not explicitly consider state aid, methods for incorporating cost indices such as ours into state aid formulas are well known.³ The main contribution of our approach is the development of new methods to select educational outcome measures and to control for school district efficiency. Moreover, the application to New York is instructive because school districts in the state have a wide variety of educational environments, from sparsely populated rural areas to large central cities.

Our approach is consistent with many of the principles underlying recent educational reform efforts. In particular, many states have moved from process-oriented to outcome-oriented policies,

such as the development of common standards and achievement measures. Moreover, many states have implemented programs designed to encourage school choice and efficiency.⁴ Despite this new focus, recent reform efforts have not recognized, for the most part, that outcomes and efficiency cannot be accurately compared across districts without a viable method for measuring educational costs. Some reforms, including those in South Carolina and Dallas, have discovered that performance measures will be worse, on average, for low-income than for high-income schools and make ad hoc adjustments to account for this cost-related effect.⁵ However, these reforms do not explicitly recognize the role of input costs or environmental factors, and their adjustments do not accurately account for cost variation across schools or school districts.

Our analysis shows how to estimate cost differences across districts controlling for district efficiency, but a complete analysis of the role of cost indices in state aid formulas is beyond the scope of this chapter, largely because district efficiency may be influenced by state aid. Moreover, as many other chapters in this book make clear, education reform requires changes in incentive systems and school management as well as in school finance. Our objective in this chapter is to highlight the importance of educational cost differences across districts so that these differences can be incorporated into broad school reform efforts. The cost models we develop can incorporate the new performance measures which have been developed in recent education reforms.

Educational Production and Costs

Our approach builds on the large literature on educational production functions and educational costs. This section reviews the key elements from this literature and discusses the unique features of our approach. The following section presents our empirical analysis.

Educational Production Functions

The literature on the technology of public education focuses on a production function of the form:⁶

$$(1) \quad S_{it} = \alpha I_{it} + \beta X_{it} + \delta S_{i,t-1} + e_i + \mu_{it} .$$

The subscripts i and t indicate school and time, respectively; S is a measure of educational service or output, such as a test score or a drop-out rate; I is a vector of inputs, such as teachers and classrooms; X is a vector of environmental factors, such as the share of students with learning disabilities; e is a set of unobserved characteristics of the school and its pupils; μ is a random error term; and α , β , and δ are parameters to be estimated.⁷ The lagged value of S captures the continuing impact of inputs, environmental factors, and random errors in previous years on this year's output; its coefficient, δ , measures how fast the output from the previous year "deteriorates" between school years.⁸

Environmental factors, X , also have been called "external" inputs, that is, inputs not controlled by school officials. The term "environmental factors" is taken from the literature on local

public finance,⁹ whereas the term "external inputs" is taken from the literature on school production functions. Although these two literatures developed separately, these two terms refer to exactly the same concept. Several recent studies have brought these two strands of literature together.¹⁰

If observations for each school are available at three points in time, this equation can be transformed into change form:

$$(2) \quad S_{it} - S_{i,t-1} = \alpha (I_{it} - I_{i,t-1}) + \beta (X_{it} - X_{i,t-1}) + \delta (S_{i,t-1} - S_{i,t-2}) + (\mu_{it} - \mu_{i,t-1}).$$

In this case, the dependent and explanatory variables are expressed in change form and the school-specific effect, e_i , cancels out. Without differencing, the unobserved school-specific effect can be a source of omitted variable bias in equation 1.¹¹

Focusing on a specific educational output is the most direct way to look at the technology of public education. Moreover, this approach can be applied to school districts, schools, classrooms, or even students.¹² The more micro levels of focus make it possible to isolate the variables that influence the interaction between students and teachers that is at the heart of this technology.

This approach also has some disadvantages, however. The principal one for our purposes is that it focuses on one output at a time.¹³ Public schools are complex institutions that provide many different outputs, which are likely to share inputs and influence each other, that is, to be produced jointly.¹⁴ As a result, it is difficult to make statements about the technology of production

for all educational outputs on the basis of equations 1 or 2. This limitation is crucial for us because our objective is to determine the differences in technology, and the associated differences in costs, for the unit that is evaluated and aided by state government, namely the school district. We need an overview of educational technology in a district, not the specific classroom technology for a single educational output.

Educational Cost Functions

This problem leads us to the principal alternative method for studying educational technology, namely an analysis of school spending or costs, defined as the sum of input purchases.¹⁵

Associated with every production function, such as equation 1, is a cost function. However, cost functions are only observed at the district level, in effect after the cost functions for various educational outputs have been aggregated. Let the subscript j indicate the school district. Suppose S^* is an index of educational output for a school district, E is spending, and AC indicates expenditure per unit of S_{jt-1}^* . Spending is measured in per pupil terms. Then by definition:

$$(3) \quad E_{jt} = (S_{jt}^*) (AC_{jt}) .$$

Moreover, a general form of the average cost function is (ignoring past history, S_{jt-1} , for the moment)

$$(4) \quad AC_{jt} = c (S_{jt}^* , P_{jt} , X_{jt} , \epsilon_j , v_{jt}) ,$$

where P is a vector of input prices, e is a set of unobserved district-specific variables, and v is a random error term. Combining equation 3 and equation 4 yields

$$(5) \quad E_{jt} = h (S_{jt}^* , P_{jt} , X_{jt} , \varepsilon_j , v_{jt}) .$$

Before estimating equation 5, we must deal with three major conceptual issues. The first issue is that S^* clearly is endogenous; school districts make spending and service quality decisions simultaneously.¹⁶ Fortunately, the literature on the demand for public education provides extensive instruments to use in a simultaneous equations procedure for equation 5. In particular, the standard median voter model of education demand shows how public service quality, S^* in our approach, depends on income, intergovernmental aid, tax-share (usually specified as the ratio of median to mean property values), and preferences.¹⁷

As an aside, this approach is based on the auxiliary equation

$$(6) \quad S_{jt}^* = d (D_{jt} , \psi_{jt}) ,$$

where D is a vector of demand variables and ψ is a random error term. This equation can be substituted into equation 5 to provide an alternative to our basic approach. Ultimately we will compare cost indices based on equation 5 with cost indices based on equation 5 after equation 6 has been substituted into it. However, this "reduced form" approach has a major disadvantage

compared to estimating the structural equation 5, namely that environmental factors influence a voter's tax price, which is her tax share multiplied by the marginal cost of public services, and therefore are demand factors themselves. Hence, the coefficients of environmental variables in the reduced-form approach reflect both their direct impact on educational costs, which is the effect we are after, and their indirect impact through demand. These effects cannot be untangled without assuming specific functional forms for the relationships in the model forms that cannot be tested.¹⁸ We prefer the structural approach because it requires no such assumptions.

The second issue concerns how to measure S^* . One possible approach is to include every possible measure of school outputs and let the regression procedure determine how they are weighted to form S^* . This approach has two serious problems. First, because output measures often are highly correlated with each other, it introduces extensive collinearity into the regression. This collinearity may make it impossible to estimate any coefficients with precision, including the coefficients of the cost variables. Second, this approach undermines our ability to identify the model, since every new (endogenous) output measure requires another instrument. As a practical matter, therefore, the key problem is how to pare down the set of school output measures in a sensible way.

Equation 6 provides a partial solution to this problem. When analyzing district spending, one is interested in school output measures that households care about, as reflected in their demand

for public services. School output measures that are uncorrelated with demand variables do not fit the bill.

This is only a partial solution to the problem, however, because equation 6 includes an error term, ϵ , so that some outputs people care about at a district level may not be correlated with demand variables, at least not with demand variables we can observe. Hence, evidence that an output variable is correlated with demand variables must be combined with judgements about the importance of various output measures based on previous literature. Our judgements on this issue are presented below.

The third issue is that equation 5 includes two error terms, which we do not observe and which might be a source of bias. The district-specific effect, e , which captures all unobserved variables that do not vary over time, can be eliminated through differencing, as in equation 2.¹⁹ For our purposes, however, differencing has two serious limitations. First, this procedure makes it impossible to observe the impacts of input and environmental factors that do not vary over time; these impacts are buried in the district-specific effect. Because many input and environmental factors vary only slowly over time (and often cannot even be observed every year), this approach may mask most of the variation in costs across districts and is not appropriate when one's objective is to obtain a comprehensive cost index.

Second, as seen in equation 2, differencing does not eliminate variables that vary over time. These variables are likely to include many inputs and environmental factors, along with unobserved

outputs, a district's past service quality, and its degree of inefficiency.²⁰ Unless these variables are accounted for, estimated coefficients for input and environmental factors that are included are likely to be biased--even with differencing. Moreover, this bias could be upward or downward, depending on the correlation between the included and excluded variables.

To deal with these problems, we estimate the undifferenced form of the cost function with a new control variable designed to capture all the systematic components of both e and v . The variable we use is based on a technique called Data Envelopment Analysis, or DEA, which has been used to measure school district inefficiency²¹. Cost inefficiency is the extent to which a district is spending more than necessary to obtain its output level. This inefficiency consists both of using too many inputs for a given amount of output (technical inefficiency) and of using the wrong combination of inputs given input prices (input allocative inefficiency). Further explanation of DEA is provided in the appendix.

As it turns out, a standard DEA "efficiency" measure captures the impact of any factor that influences the relationship between service quality and costs--not just district efficiency. All else equal, an efficient district can obtain the same service quality as an inefficient district at a lower cost. As discussed in the appendix, however, the relationship between service quality and costs is also affected by environmental factors.²² Consider two equally efficient districts, one of which has a very harsh environment compared to the other. The district with the harsher environment will have to spend more to obtain the same service quality. Hence a standard DEA measure picks up the

impact of environmental factors as well as of efficiency. The same logic applies to any other unobserved systematic factor in either error term of equation 5. Districts that made relatively high investments in education in the past, for example, will have a favorable legacy that allows them to obtain the same service quality as other districts at a lower cost.²³

Thus, including a standard DEA "efficiency" measure will eliminate the potential bias from the unobserved, and hence omitted, non-cost variables included in the two error terms in equation 5.²⁴ DEA captures the impact of any factor that influences the relationship between service quality and costs, so our DEA variable is a comprehensive insurance policy against omitted variable bias.

Unfortunately, however, this insurance policy has a price, namely the resulting duplication of contemporaneous input and environmental cost variables.²⁵ To be specific, input prices and environmental factors are included as the X s in equation 5 as well as in the DEA variable. As a result, some of the full impacts of input prices and environmental factors on costs will be captured by the estimated coefficients of the X s and some will be captured by the DEA variable's coefficient. We do not know exactly how these impacts will be divided, but we do know that the true impacts will not be fully captured by the X s, that is, by the observed values of the input prices and environmental factors.

Our cost indices are based solely on the coefficients of the X s and are not affected by the coefficient of the DEA control variable. It follows that our approach inevitably provides an

underestimate of the impact of input prices and environmental factors on costs; some of the true impact of the environment is buried in the DEA coefficient.

In addition, our approach focuses on the role of contemporaneous input prices and environmental factors and ignores past values of these variables. One could argue that a cost index should capture past as well as current values of these variables. A district should be compensated, the argument might go, for the lingering effects of a relatively harsh environment in the past, as well as for a harsh environment in the present. This argument has some appeal, but it also raises many unresolved issues, such as how far back in history to go. Moreover, past history is difficult to observe and incorporate into the model. To the extent that contemporaneous values of input prices and environmental factors are highly correlated with past values, our approach may pick up some past history. But neither our approach nor any previous research produces cost indices that include a comprehensive treatment of each district's history of input prices and environmental variables.

In short, our approach provides a conservative estimate of the impact of contemporaneous input prices and environmental factors on school district costs. Although an exact cost index would be preferable, no procedure for estimating such an index is yet available, and our approach has the advantage that the estimated coefficients are not biased upward (in absolute value) because of unobserved district inefficiency or past effort. Moreover, a focus on contemporaneous, as opposed to past, input prices and environmental factors, is appropriate given the complex role (philosophically and technically) of past history and the limitations on available data.

One final point: A DEA "efficiency" measure also might be endogenous; that is, some of the same factors that influence decisions about spending might also influence decisions that lead districts to act in an efficient manner. To account for this possibility we identify an instrument for district efficiency and treat the DEA measure as endogenous.

Cost Indices

For the purposes of designing intergovernmental aid formulas, one needs a measure of the cost, based on factors outside a district's control, of providing a given quality of education.²⁶ Educational quality is defined by the educational outputs, S . Because equation 5 determines the impact of input and environmental costs on spending **holding S constant**, it is ideally suited for calculating a cost index. This approach has been applied both for school and non-school spending.²⁷ Our cost indices are calculated in the same way as the indices in previous studies; as explained below, these cost indices use the estimated regression coefficients to calculate the amount each district would have to spend to obtain average quality public services.

An alternative approach to cost indices based on compensating wage differentials also has appeared in the educational literature. According to this approach, some districts have to pay higher wages than other districts to attract teachers of the same quality. Several studies have estimated the extent to which teacher wages vary across districts based on factors outside a district's control (accounting for factors that a district can control) and then calculated a wage index based on this estimation.²⁸

The problem with this approach is that it dramatically minimizes the role of the school environment. A comprehensive cost index needs to account not only for the fact that some districts must pay more than others to hire teachers of any given quality, but also for the fact that some districts must hire more teachers than others to provide the same quality educational outputs for their students. Indices based on wages alone therefore inevitably provide an incomplete and potentially misleading picture of cost variation across districts.²⁹ We will demonstrate this problem using our New York data.

Empirical Analysis of Costs in New York School Districts

We estimate cost models and education cost indices for 631 school districts in New York in 1991.³⁰ This section describes our measures, data sources, and empirical analysis of education costs, and it provides a comparison of alternative education cost indices.

Measures and Data Sources

Table 1 provides descriptive statistics for the variables used in the analysis. A district's approved operating expenses (AOE) per pupil, which is provided by the New York State Department of Education, is used to measure expenditure. AOE includes salaries and fringe benefits of teachers and other school staff, other instructional expenditure, and all other expenditure related to operation and maintenance of schools.³¹ Average AOE per pupil for the sample was about \$6,054.

Potential school outcome measures in our data range from standardized test scores to dropout and graduation rates. Both the production and cost literature have relied most heavily on average achievement test scores as output measures.³² A few studies also have emphasized the role of test score distributions.³³ One argument in favor of distributional measures, such as standard deviations, is that education to some degree serves a screening function. As one scholar points out, "In a screening model, the output of schools is information about the *relative* abilities of students."³⁴ This would suggest that more attention should be directed toward the distribution of observed educational outcomes (instead of simply the means)." Several studies also have focused on the high school drop-out rate.

As discussed previously, collinearity severely limits the number of outcomes that can be included in a cost model. We used a three-step process to select a reduced set of outcome measures. First, we identified outcomes that appear to be related to voters' willingness to pay for education by regressing each potential outcome measure on a set of education demand variables, including income and tax share. Using a broad definition of "related," namely an adjusted R-squared of at least 0.1, we were able to eliminate both the average and the standard deviation of standardized achievement test scores as outcome variables for this analysis.³⁵

Second, from the set of outcomes correlated with demand factors, we identified subsets of variables that, based on previous research, appeared to be reasonable measures and then, where appropriate, calculated an average across the variables in such a subset. This step led to three

outcome measures, all of which capture the tails of the student achievement distribution, instead of the average as in much previous research. The first of these measures is based on Pupil Evaluation Program, PEP, tests given to all third- and sixth-grade students in reading and math. The specific measure is the average percentage of students performing above a standard reference point on these four exams. The standard reference point is used to identify students requiring special assistance (and Chapter 1 funding from the federal government). The second measure is the percentage of students receiving a special Regents diploma upon graduation from high school. Regents diplomas are given to students who pass standardized tests given by the state to high school students. To balance this measure of achievement, the third measure is the inverse of the drop-out rate, namely the percentage of students not dropping out of school.³⁶

Third, we used factor analysis to determine whether the selection and clustering of our outcome measures adequately captured the variation in the data across all potential such measures. The size and pattern of the factor scores strongly supports our choices.³⁷

As explained earlier, a cost model should control for unobserved district characteristics that influence costs. Using the DEA method discussed in the previous section, a standard cost "efficiency" index was constructed for each school district based on AOE per pupil and the three outcome measures presented in Table 1. As explained earlier, this index captures not only efficiency but also environmental cost factors and past school decisions that shift the cost frontier facing a school district. Because this index is held constant in constructing the cost indices, we are

being conservative in our estimate of costs; that is, our cost indices ignore any cost effects picked up by the DEA index instead of by the input and environmental variables in the cost model. The average "efficiency" score is 0.66, with 23 districts (4 percent) with an index of one and 350 districts (55 percent) with an index below 0.7.

Cost differences across districts reflect both input price differences and environmental factors. To measure input price differences, we estimated a teacher salary index. This index adjusts for differences in teacher experience, education, and certification to reflect differences in the cost of teachers of equivalent quality.³⁸ A potential problem with the index is endogeneity arising out of the relationship between teacher salaries and spending decisions.³⁹ It is possible that some of the variation in teacher salaries reflects discretionary decisions by district administrators, not underlying differences in opportunity wages for teachers. To avoid this problem the index is based on salaries of teachers with five years or less of experience. Even if excessive expenditures are used primarily to increase teacher's salaries, this benefit is less likely to accrue to the most recently hired teachers. Moreover, as explained below, this wage variable is treated as endogenous.

For the most part, the cost literature focuses on one environmental variable, namely the number of students. The central question addressed in this literature is whether per-pupil costs rise or fall when the number of pupils increase, that is, whether there are economies to pupil scale.⁴⁰ Because many studies find that expenditures per pupil are a U-shaped function of enrollment, we include enrollment and its square as environmental variables.⁴¹ Past studies have also considered

the share of students in secondary grades, the share of students in special education programs, the share of students with limited English proficiency, and the share of students receiving a subsidized lunch.⁴²

The education production literature has highlighted the importance of family background and student characteristics.⁴³ Our data set allows us to measure several environmental variables in these categories, namely the percentage of children in poverty, the percentage of households with a female single parent, the percentage of children with limited English proficiency, the percentage of students with a handicapping condition, and the percentage of total enrollment that is high school students.⁴⁴

Service outcomes, the efficiency index, and the price of labor are all determined simultaneously with district spending through discretionary decisions made in the annual budgeting process. To control for this endogeneity, our cost model is estimated using two-stage least squares, with an appropriate set of additional instruments. The instruments associated with the service outcomes are drawn from the literature on the demand for public services.⁴⁵ Following a standard median voter model, we use median income as a fundamental determinant of voter demand. Demand also depends on intergovernmental aid; our state aid variable, basic operating aid, is the principal form of non-categorical aid provided to school districts in New York.⁴⁶ The standard tax price facing the median voter equals her tax share multiplied by the marginal cost of educational services. The marginal cost component is already in the cost model (in the form of the input price

and environmental factors), but the tax share makes a suitable instrument. We measure the tax share with the ratio of median to mean residential property value and with an estimate of the district's ability to export some commercial and industrial property taxes onto non-residents.⁴⁷ Finally, we include several socio-economic variables that are likely to be related to demand for education, namely the percentage of households with children, the percentage of households living in owner-occupied housing, and the percentage of adults with a college degree.⁴⁸

We also use instruments associated with the price of labor or the efficiency index. Since comparable private sector prices for teachers were not available, we use 1990 county population as a instrument for teacher salaries. Our choice of this instrument is based on the stylized fact (and a central prediction of urban economics) that the cost of living, and hence, the cost of hiring workers, increases with metropolitan population. Identifying instruments for the efficiency index is more difficult. While there is a large literature on bureaucratic behavior, there is little associated empirical literature examining the causes of inefficiency.⁴⁹ The bureaucratic models suggest that greater inefficiency will be associated with larger and wealthier school districts, those facing less competition, and those with poorer performance incentives for their employees. Enrollment and median income already have been included as exogenous variables. Good measures of private school competition are not available, but competition also may come in the form of voter referenda on school budgets. In New York, all school districts are required to have budget referenda except

for city school districts, where the budget is set entirely by elected city officials. A dummy variable for city districts therefore is included as an instrument for the efficiency index.⁵⁰

Cost Model Results

We estimate our education cost models using a modified Cobb-Douglas cost model with a quadratic enrollment term. The Cobb-Douglas form imposes several restrictions on the production technology for educational services.⁵¹ The simplicity and conceptual plausibility of this function along with its frequent successful application in empirical research outweigh its potential limitations.⁵²

The dependent variable is the log of AOE per pupil. The cost models were estimated using linear 2SLS, with outcome measures, the efficiency index, and the price of labor treated as endogenous.

Our initial specification, called Model 1, is presented in the first column of Table 2. This specification, which is based on the three outcome measures defined above, performs very well. The outcome measures all have positive coefficients, as expected, and two of the three coefficients are statistically significant. The PEP test scores variable (the average percentage of students above standard reference point) has a *t*-statistic of 1.5. The *efficiency@* index has, as expected, a negative coefficient and is statistically significant; greater efficiency in a school district is associated with lower expenditure, *ceteris paribus*.

Moreover, six of the eight cost variables have statistically significant coefficient with the expected signs. The teacher salary variable is, as expected, positively related to expenditure and its coefficient is quite large; in fact, a 1.0 percent increase in teacher salaries is associated a 0.89

percent increase in per pupil expenditure. Both enrollment variables are statistically significant and indicate a U-shaped per pupil expenditure function. Based on these results, the "minimum cost enrollment" falls at a district enrollment of about 3,300 pupils.⁵³ Child poverty rates and the percentage of female headed households, included to reflect family background, are both positively related to expenditure and statistically significant, and we find a positive and significant relationship between spending and the share of high school students with limited English proficiency. The percentage handicapped and percentage high school variables also have the expected signs but their t-statistics are just above 1.0. Overall, this regression provides strong confirmation of our approach; by controlling for (endogenous) outcome measures, efficiency, and past history, one can precisely measure the impact of many contemporaneous input and environmental cost variables on school district spending.

We also estimated several variants of this model to determine the robustness of our results.

In Model 2 we explore one possible explanation for the insignificance of the percentage handicapped variable, namely the heterogeneity of the students in this category and the associated variation in the special services they need. Using disaggregated information on handicapped students in New York by the level of service then receive, we examined several handicapped variables in the cost model.⁵⁴ The percent of students with severe handicapping conditions (requiring special services out of the regular classroom at least 60 percent of the school day) does have a statistically significant positive affect on district expenditures. A one percentage point

increase in these students raises per pupil expenditures by close to one percent. The other outcome and cost factors remain statistically significant with little change in their coefficients. Model 2 is our preferred specification and is used to construct our principal cost index in Tables 3 and 4.

Because one of our outcome measures is not statistically significant, we also estimated cost models using two different pairs of outcome measures. The resulting models, called Models 3 and 4 in Table 2, each include a DEA efficiency index based on only the two outcome measures in the model. In both cases, the coefficient of the PEP scores variable is statistically significant with a magnitude similar to that in Model 1. These results reinforce the importance of controlling for elementary student performance in the construction of cost indices and suggest that it may be collinearity that keeps down the significance of the PEP variable in Models 1 and 2. Because it provides a broader range of outcome measures, we will utilize the three-outcome model to construct our education cost indices.

Comparison of Education Cost Indices

The cost models in Table 2 can be used to construct comprehensive educational cost indices. Our cost index is designed to capture the key cost factors outside of a district's control, including the underlying cost of hiring teachers (the opportunity wage), district size, family background, and student characteristics. Variation in expenditure among districts that reflects differences in service quality, in efficiency, or in past history is eliminated from the calculations; that is, service quality and efficiency are held constant across districts. To be specific, we multiply

regression coefficients by actual district values for each cost factor (and by the state average for outcomes and efficiency) to construct a measure of the expenditure each district must make to provide average quality services given average inefficiency.⁵⁵ Our cost indices simply express this predicted expenditure relative to the state average.⁵⁶

The first column of Table 3 presents our principal cost index, which is based on Model 2 in Table 2. This index has a range from 78 to 240 with a standard deviation of 17. Seventy-five percent of the districts have indices below 105, and 75 percent have indices above 90.

Table 3 also presents several alternative cost indices. Columns 2 and 3 presents cost indices based on alternative cost models; the cost model in column 2 has no control for district efficiency, and the one in column 3 treats district efficiency as exogenous. These columns reveal that, compared to our preferred model, ignoring efficiency tends to magnify cost differences across districts whereas treating efficiency as exogenous tends to dampen them. Because our efficiency index reflects cost factors to some degree, leaving out this index boosts the impact of the cost factors in the equation. Because the index also reflects other factors, such as efficiency, that may be correlated with costs, the index in column 2 may be affected by omitted variable bias and may therefore overstate cost differences across districts. Treating efficiency as exogenous introduces another possible bias, namely endogeneity bias. As it turns out, the effect of leaving out the efficiency variable altogether is smaller than treating efficiency as exogenous, at least on average, so the correlation between the indices in the first two columns, 0.94, is higher than

the correlation between the indices in columns one and three, 0.84. This result indicates that a cost index correcting for efficiency, which is difficult to obtain, is roughly proportional to a cost index without an efficiency correction. However, the actual distribution of aid using these two cost indices be quite different because the efficiency correction lowers variation in costs.

Table 3 also compares our preferred cost index with a cost index based on an alternative approach in the education literature and with two forms of cost indices widely used in practice. As explained earlier, if demand variables are substituted for service outcomes, then an indirect (or reduced-form) expenditure model can be used to construct a cost index.⁵⁷

Most states use some form of weighted pupil measure in the allocation of aid. In New York, for example, students with special needs, handicapping conditions, or in secondary school receive heavier weights in the distribution of aid. By taking the ratio of weighted pupils (specifically, total weighted pupil units, TWPU) to total enrollment we construct a cost index that indicates the level of cost adjustment in a typical state aid formula. This approach makes ad hoc adjustments for cost differences across some types of students and is likely to understate overall cost differences because it focuses on only a few cost-related student characteristics.

The most common cost index proposed in education research focuses on the relationship between socio-economic factors and teacher salaries. Teachers are expected to command higher salaries if they are of higher quality (or have characteristics rewarded in union contracts), or if they have to work under more adverse working conditions. Working conditions can be affected by

district decisions concerning resource utilization (pupil-teacher ratios) or by socio-economic factors out of the district's control that reflect the harshness of the education environment (such as a relatively high incidence of special needs or disadvantaged children). By holding teacher quality, demand variables, and discretionary resource factors constant, these studies have constructed education cost indices to reflect the wage differentials required to "compensate" for an adverse socio-economic environment.⁵⁸ While a compensating wage-based cost index may capture cost factors associated with higher teacher salaries, it does not control for differences across districts in resource usage (including hiring of teachers!) required to provide a given level of service outcomes.

The last three columns of Table 3 present these alternative education cost indices. The indirect cost index, which does not control for inefficiency has slightly lower variability than our preferred cost index in column 1.⁵⁹ The least variability appears in the weighted-pupil and teacher-salary indices, largely because these indices are only capturing a portion of actual cost differentials.

Correlation coefficients reiterate the substantial differences among these indices. The correlation between our preferred index and the indirect index is 0.63, which suggests that the indirect approach may not do a good job controlling for service quality differences and may therefore result in biased cost indices.⁶⁰ The correlation between our preferred index and the weighted-pupil index is extremely low, only 0.14; the approach used by New York State therefore misses most of the actual variation in costs across districts. Finally, the correlation between our preferred index and the teacher salary index is 0.47, indicating only a moderate correlation between

the factors that push up the salary needed to attract a given quality of teacher and the factors that push up the cost of providing a given quality of educational services. The teacher salary index is not related to either the indirect cost index or the weighted pupil index.

To provide a more disaggregated comparison of these cost indices, Table 4 presents average index scores by region, enrollment size, and income and property wealth of school districts.

The direct cost index with endogenous efficiency identifies the large upstate central cities and downstate small cities as having the highest costs. (The large downstate cities, New York City and Yonkers, are not included in sample due to missing data.) This result reflects higher teacher salaries in downstate districts and higher environmental cost factors in upstate cities. Upstate suburbs and rural districts have below average costs. This table also clearly shows the U-shaped relationship between costs and enrollment and reveals that costs tend to be slightly higher for both the poorest and the richest districts, measured by either income or property wealth. Higher income or wealth districts, particularly in downstate New York, may have a relatively favorable educational environment, but they must pay relatively high teacher salaries.

Table 4 also shows the values for the alternative indices in each of these categories. Compared to our preferred index, the cost differences across types of district are magnified somewhat with the no-efficiency index and dampened considerably with the weighted-pupil, teacher-cost and indirect cost indices. Comparing the various indices by pupil-size category reinforces the similarity between our preferred index and the no-efficiency index, but also reveals

substantial differences between our preferred index and the others. The indirect cost index accentuates the U-shaped relationship between enrollment and per pupil costs, while the other indices understate this relationship. In general, they completely fail pick up the relatively high costs of small districts and understate the costs of the largest districts.⁶¹ Comparisons based on income class or property value class also identify several distinct differences between indices. While our preferred index shows little variation across income (and property wealth) classes, the no-efficiency, indirect cost, and teacher salary indices show substantially higher costs in low-income districts. These differences are difficult to interpret since they could reflect either inefficiency or unobserved environmental cost factors.

What types of districts tend to have particularly high or low costs and which environmental factors principally account for these cost differences? To answer this question we examined the ten percent of school districts with the highest and lowest costs (Table 5). Average values for environmental factors for these districts are compared to the state average. For high-cost districts, costs average 52.7 percent above the state average, \$3,046 per pupil. All upstate large cities and over 70 percent of downstate small cities qualify as high-cost districts. Over 10 percent of downstate suburbs and upstate small cities also fall in this category. Enrollment, percent of children in poverty and with limited English proficiency, and percent of single-parent female-headed households are all well above the state average in these districts.

Combining the environmental indices with the regression coefficients for model 2 in Table 2, we can identify which environmental factors have a particularly strong effect on costs. Higher teacher salaries and a relatively high number of female-headed households each account for over 30 percent of the higher costs in these districts. Limited English proficiency and poverty are also important factors driving up costs. The higher enrollments in some high-cost districts may actually lower per pupil costs, because their enrollments are close, on average, to the cost minimizing enrollment.

The 10 percent of districts with the lowest costs have costs 20 percent below average, \$1,091 per pupil. Most of these districts are upstate suburbs; a few are rural districts. Poverty, female-headed households, severely handicapped students, and students with limited English proficiency are all relatively uncommon. Lower teacher salaries, lower poverty rates, fewer female-headed households, and higher enrollments each account for 20 percent of the lower costs in these districts.

Conclusions and Policy Implications

At the conceptual level, the importance of educational costs cannot be denied. Through no fault of their own, some school districts must spend more than other districts to obtain the same level of educational outcomes. Despite widespread agreement on this point among scholars, educational cost indices remain illusive because any method to estimate them must overcome complex methodological obstacles. Given the stakes involved, namely the allocation of state educational aid, we believe that overcoming these obstacles is one of the principal challenges facing scholars and policy makers interested in education finance. This paper develops and implements a method for estimating educational cost indices that resolves some of these difficulties.

Our approach, like several others, focuses on the impact of input prices and environmental cost factors on educational spending, controlling for educational service quality. This approach leads to an index of the amount a school district would have to spend, given the input prices and environment it faces, to obtain average-quality educational services. Our contributions are to develop new criteria for selecting service quality measures and to explicitly control for school district efficiency and other unobserved district characteristics that might lead to biased cost indices.

When applied to data for school districts in New York state, our approach works well in the sense that most of the regression coefficients are statistically significant and all of them have the anticipated signs. Hence, the cost indices we estimate control for a variety of service quality measures (as well as district efficiency) and estimate with precision the impact of input prices and

environmental factors on educational costs. The major disadvantage of our approach is that it requires the calculation of a complex "efficiency" measure, based on Data Envelopment Analysis. This disadvantage may make our approach impractical as a tool for designing school aid formulas.

We also find, however, that cost indices based on a cost model that does include the DEA index are highly correlated with those based on our preferred cost model. Thus, school aid formulas based on this simpler formula might be acceptable. However, a better compromise would be to discover simpler methods to control for district efficiency and other unobserved district characteristics and to include these methods in a cost model. We also find two widely used methods for estimating educational costs, namely those based on weighted pupils and on required teacher salaries, do not provide reasonable approximations for our method, which is to be preferred on conceptual grounds. The weighted-pupil cost index used in New York is virtually uncorrelated with our cost index, and the teacher-salary index is only moderately correlated, misses the U-shaped relationship between costs and enrollment, and greatly understates the costs in large city districts. In our judgement, therefore, these approaches are seriously deficient.

Educational cost variation across school districts is a crucial issue that has not been adequately recognized by either courts or state legislatures. Despite its fundamental consistency with a focus on school performance, it also has not been adequately incorporated into recent performance-based school reform efforts. The large literature on production and cost in education provides a solid foundation for the development of education cost indices. This paper demonstrates

the serious flaws in existing ad hoc indices, which do not build on this foundation, and shows how more acceptable cost indices can be derived.

APPENDIX

Measuring Inefficiency in Public Services

Several methods for estimating technical and cost efficiency have been developed over the last several decades. The non-parametric method used in this paper, Data Envelopment Analysis, DEA, is based on production theory in economics and has been operationalized as DEA since the late 1970s.⁶² One of the major advantages of DEA is that it is non-parametric, that is, it requires no *a priori* specification of the functional form. One disadvantage is that the technique is non-stochastic.⁶³

These methods have been extended to analyze costs and economies of scope in public sector production. The relevant mathematical programs are solved to compare the expenditure of a given local government with the expenditure of other local governments producing the same level of services. If the local government is producing at the cost-minimizing level, then no other local government (or linear combination of local governments) is producing the same level of services with lower expenditure.⁶⁴

One problem with existing DEA methods for estimating inefficiency is the maintained assumption that the technology can be represented by one frontier. This assumption presumes that all deviations from the cost frontier are attributable to inefficiency. While DEA has been commonly employed to examine public organizations such as school districts, the assumption of one cost frontier is not consistent with the nature of public production.⁶⁵ As explained in the text, input

prices, P , and exogenous socio-economic variables, X can have an important influence on the translation of government activities into service outcomes. As a result, there will be multiple cost frontiers reflecting differences in P and X . Estimates of the minimum level of costs and cost inefficiency that do not control for these cost factors will be biased.

Recently, a method has been developed for estimating technical and cost efficiency that allows for multiple frontiers.⁶⁶ Figure 1 illustrates two minimum cost frontiers assuming for simplicity one service outcome, S . For all levels of S , $C(S/P_1, X_0) \geq C(S/P_0, X_0)$ because $P_1 > P_0$. Efficiency estimates should be made in reference to the correct frontier. A local government is said to be *cost efficient* if the observed level of expenditure is equal to the minimum total cost of providing the observed level of services, given resource prices and environmental conditions.

While this method provides a more realistic estimate of relative cost efficiency among school districts, it can handle only a few fixed cost factors, and these fixed cost factors must be selected prior to estimation of the cost model. Selected cost factors may turn out to be statistically insignificant, so that a complex iterative procedure would have to be developed to make the regression and the DEA consistent. To avoid these problems, we use the unadjusted cost efficiency index, which compares all districts to the cost frontier for the district with the most favorable environment. Specifically, our measure of cost "efficiency," θ , is equal to C/E , where C equals minimum costs and E equals actual expenditure.

If local governments are cost efficient and face the most favorable cost environment, then expenditure reflects the minimum cost of providing services and θ equals 1.0. In any other case,

that is, with either inefficiency or unfavorable fixed factors, θ is less than 1.0. To illustrate, assume P_0 and X_0 in Figure 1 represent the most favorable educational environment (minimum cost frontier for district D). The cost efficiency index for district H would be $\theta_H = C(S/P_0, X_0)/E_H$. Since district H faces higher factor prices, $C(S/P_1, X_0)/E_H$ represents the true (unobserved) cost efficiency and $C(S/P_1, X_0)/C(S/P_0, X_0)$ the index of environmental harshness.

Endnotes

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1. For a good discussion of past court challenges see Allan Odden and Lawrence Picus, *School Finance: A Policy Perspective* (McGraw-Hill, Inc., 1992). Steven Gold, David Smith, Stephen Lawton, and Andrea C. Hyary, *Public School Finance Programs of the United States and Canada, 1990-91* (The Nelson A. Rockefeller Institute of Government, 1992) provide a good overview of state aid systems in the early 1990s.

 2. See Helen Ladd and John Yinger, "The Case for Equalizing Aid," *National Tax Journal*, 47 (March 1994), pp. 211-224.

 3. See Ladd and Yinger, "The Case for Equalizing Aid," pp. 211-224, for an overview of adding costs to several types of aid formulas.

 4. For general discussion of education reform and school choice, see Susan Fuhrman, Richard Elmore, and Diane Massell, "School Reform in the United States: Putting it into Context," in S. Jacobson and R. Berne, eds., *Reforming Education: The Emerging Systemic Approach* (Thousand Oaks, CA: Corwin Press, Inc., 1993), Bruce Cooper, "Educational Choice: Competing Models and Meanings," in S. Jacobson and R. Berne, eds., *Reforming Education: The Emerging Systemic Approach* (Thousand Oaks, CA: Corwin Press, Inc., 1993), and Eric Hanushek, *Making Schools Work: Improving Performance and Controlling Costs* (The Brookings Institution, 1994).

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5. See C. Clotfelter and Helen Ladd, "Picking Winners: Recognition and Reward Programs for Public Schools," this volume.
 6. This literature is reviewed in Eric Hanushek, "The Economics of Schooling: Production and Efficiency in Public Schools," *Journal of Economic Literature* 24 (1986), pp. 1141-1177, Elchanan Cohen and Terry Geske, *The Economics of Education* (Pergamon Press, 3rd edition, 1990), and David Monk, *Educational Finance: An Economic Approach* (McGraw Publishing Company, 1990).
 7. This error structure assumes that the error component, e , does not vary over time. This assumption may not be appropriate if the error component includes student cohort effects as well as school effects. For simplicity, this equation is written in linear form, although other forms can be used.
 8. The concept of service "deterioration" is discussed at length in the chapter by Ronald Ferguson and Helen Ladd, "Additional Evidence on How and Why Money Matters: A Production Function Analysis of Alabama Schools," this volume. Note that the specification with a lagged value of S also can be derived by including lagged values of I , X , and μ and then subtracting the equations for two succeeding years. With this approach, the coefficients for the lagged values of I and X must "deteriorate" over time at the same rate such that $\beta_{t-i} = d^i \beta_t$, where β_t is the coefficient vector for these variables in year t .

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9. David Bradford, Robert Malt and Wallace Oates, "The Rising Cost of Local Public Services: Some Evidence and Reflections," *National Tax Journal* 22 (June 1969), pp. 185-202; Katherine Bradbury, Helen Ladd, Mark Perrault, Andrew Reschovsky, and John Yinger, "State Aid to Offset Fiscal Disparities across Communities," *National Tax Journal* 37 (June 1984), pp. 151-170; and Helen Ladd and John Yinger, *America's Ailing Cities: Fiscal Health and the Design of Urban Policy* (The Johns Hopkins University Press, 1991).
 10. These studies include Kerri Ratcliffe, Bruce Riddle, and John Yinger, "The Fiscal Condition of School Districts in Nebraska: Is Small Beautiful?" *Economics of Education Review* 9 (1990), pp. 81-99; Richard Fenner, *The Effect of Equity of New York State's System of Aid for Education*, Ph.D dissertation, Syracuse University, 1991; and Thomas Downes and Thomas Pogue, "Adjusting School Aid Formulas for the Higher Cost of Educating Disadvantaged Students," *National Tax Journal* 47 (March 1994), pp. 89-110.
 11. Because this step also requires at least three years of data, it has not been taken by any production function study of which we are aware. The excellent chapter by Ferguson and Ladd, "Additional Evidence on How and Why Money Matters: A Production Function Analysis of Alabama Schools," for example, is based on one of the most complete data sets in the literature, but it estimates equation (1), not equation (2).
 12. As pointed out by Summers and Wolfe and Hanushek, among others, in applying this model to

individual students one must distinguish between individual and family background variables, peer-group variables, and school variables. See Anita Summers and Barbara Wolfe, "Do Schools Make a Difference?" *American Economic Review*, 67 (September 1977), pp. 639-652; and Eric Hanushek, "Conceptual and Empirical Issues in the Estimation of Educational Production Functions," *The Journal of Human Resources*, 14 (1979), pp. 351-388.

13. Another disadvantage is that extensive data are required.
14. Multi-product production functions have typically assumed separability between outputs so that each can be estimated in a separate equation (possibly allowing correlation across error terms by using a seemingly unrelated regression method). On the other extreme, some studies (Boardman, Davis, and Sanday) in the public sector have estimated simultaneous production functions which assume that each output simultaneously influences the other. If the production of outputs share some inputs but do not necessarily cause each other, then a production function which allows for jointness of production is the most appropriate. Recently, canonical regression has been used to estimate joint production functions. However, no one has yet shown how to use this approach to develop comprehensive educational cost indices. See Anthony Boardman, Otto Davis, and Peggy Sanday, "A Simultaneous Equations Model of the Education Process," *Journal of Public Economics*, 7 (1977), pp. 23-49; John Chizmar and Thomas Zak, "Modeling Multiple Outputs in Educational Production Functions," *American Economic Review*, 73 (May 1983), pp. 18-22;

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- Kwabena Gyimah-Brempong and Anthony Gyapong, "Characteristics of Education Production Functions: An Application of Canonical Regression Analysis," *Economics of Education Review*, 10 (1991), pp. 7-17; and John Ruggiero, "Measuring Technical Inefficiency in the Public Sector: An Analysis of Educational Production," *Review of Economics and Statistics*, forthcoming, 1995.
15. Recent examples of this approach include Downes and Pogue, "Adjusting School Aid Formulas for the Higher Cost of Educating Disadvantaged Students," pp. 89-110; Kwabena Gyimah-Brempong and Anthony Gyapong, "Elasticities of Factor Substitution in the Production of Education," *Economics of Education Review*, 11 (1992), pp. 205-217; Fenner, *The Effect of Equity of New York State's System of Aid for Education*; Scott Callan and Rexford Santerre, "The Production Characteristics of Local Public Education: A Multiple Product and Input Analysis," *Southern Economic Journal*, 57 (October 1990), pp. 468-480; Ratcliffe, Riddle, and Yinger, "The Fiscal Condition of School Districts in Nebraska: Is Small Beautiful?" Earlier studies are reviewed in Cohen and Geske, *The Economics of Education*, and Monk, *Educational Finance: An Economic Approach*.
16. Despite the obvious endogeneity of service quality, we know of only two studies of educational costs that treats service quality as endogenous, namely Baum, and Downes and Pogue. See Donald Baum, "A Simultaneous Equations Model of the Demand for and Production of Local Public Services: The Case of Education," *Public Finance Quarterly*, 14 (1986), pp. 157-78; and

Downes and Pogue, "Adjusting School Aid Formulas for the Higher Cost of Educating Disadvantaged Students," pp. 89-110.

17. For a detailed discussion and literature review on these issues, see Ladd and Yinger, *America's Ailing Cities: Fiscal Health and the Design of Urban Policy* (The Johns Hopkins University Press, 1991). The specific instruments we use are presented below.
18. This point is made by Schwab and Zampelli, and Downes and Pogue. A detailed exposition of the necessary structure is provided in Ladd and Yinger. One important assumption that is required to identify cost parameters in a reduced-form model is constant returns to scale with respect to changes in S^* . See Duncombe and Yinger for a detailed discussion of this point. See Robert Schwab and Ernest Zampelli, "Disentangling the Demand Function from the Production Function for Local Public Services: The Case of Public Safety," *Journal of Public Economics*, 33 (1987), pp. 245-260; Downes and Pogue, "Adjusting School Aid Formulas for the Higher Cost of Educating Disadvantaged Students," pp. 89-110; Ladd and Yinger, *America's Ailing Cities: Fiscal Health and the Design of Urban Policy*; William Duncombe and John Yinger, "An Analysis of Returns to Scale in Public Production, With an Application to Fire Protection," *Journal of Public Economics*, 52 (1993), pp. 49-72;
19. A first-differencing approach is used by Downes and Pogue, "Adjusting School Aid Formulas for the Higher Cost of Educating Disadvantaged Students," pp. 89-110. They are aware of the fact that

differencing eliminates some cost variables and explicitly develop cost indices only for two cost variables, namely the fractions of students receiving subsidized lunches and with limited English proficiency.

20. Hanushek points out that inefficiency may make it appear that "expenditures are unrelated to school performance." Ruggiero found evidence that inefficiency dampens the observed impact that school inputs have on outputs. See Eric Hanushek, "The Economics of Schooling: Production and Efficiency in Public Schools," pp. 1166; Ruggiero, "Measuring Technical Inefficiency in the Public Sector: An Analysis of Educational Production."
21. See, for example Shawna Grosskopf and S. Yaisawarng, "Economies of Scope in the Provision of Local Public Services," *National Tax Journal*, 43 (1990), pp. 61-74; and John Ruggiero, "Are Costs Minimized in the Public Sector? A Nonparametric Analysis of the Provision of Educational Services," Metropolitan Studies Program Occasional Paper No. 165, Center for Policy Research, The Maxwell School (Syracuse, NY: Syracuse University, 1994).
22. Ruggiero has shown how to measure inefficiency controlling for the environment in a DEA framework through the use of multiple cost frontiers (see the appendix). As explained below, however, this solution is not appropriate here. See Ruggiero, "Are Costs Minimized in the Public Sector? A Nonparametric Analysis of the Provision of Educational Services."

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23. Downes and Pogue account for past history by including both 12th-grade test scores and 11th-grade test scores for the same cohort. This is analogous to the service-quality term on the right side of equation 2, and picks up the history of cost factors, as well as of other variables. See Downes and Pogue, "Adjusting School Aid Formulas for the Higher Cost of Educating Disadvantaged Students," pp. 89-110.
 24. One important criticism of DEA is that the outputs on which it is based are selected by the researcher, not by some statistical test. This criticism does not apply to our equations because the outputs used in the DEA procedure are the same ones used in equation 5, where a statistical test of their significance is provided. See Hanushek, "The Economics of Schooling: Production and Efficiency in Public Schools," pp. 1141-1177.
 25. In principle, one could avoid this duplication by using the Ruggiero procedure to correct for environmental cost factors. However, this approach is not practical here because, as explained by Ruggiero, DEA cannot handle as many cost factors as are required for our procedure without a much larger sample of school districts than exists in any state, including New York. See Ruggiero, "Are Costs Minimized in the Public Sector? A Nonparametric Analysis of the Provision of Educational Services."
 26. For a detailed discussion of the use of cost indices in education formulas, see Ladd and Yinger,

¶The Case for Equalizing Aid, pp. 211-224.

27. For non-school spending see Bradbury et al., ¶State Aid to Offset Fiscal Disparities across Communities, pp. 151-170, and Ladd and Yinger, *America's Ailing Cities: Fiscal Health and the Design of Urban Policy*. And for school spending see Ratcliff, Riddle, and Yinger ¶The Fiscal Condition of School Districts in Nebraska: Is Small Beautiful? and Downes and Pogue, ¶Adjusting School Aid Formulas for the Higher Cost of Educating Disadvantaged Students.
28. See, for example, Jay Chambers, ¶Educational Cost Differentials and the Allocation of State Aid for Elementary and Secondary Education, *Journal of Human Resources*, 13 (1978), pp. 459-481; Jay Chambers, ¶The Development of a Cost of Education Index: Some Empirical Estimates and Policy issues, *Journal of Education Finance*, 5 (Winter 1980), pp. 262-281; Howard Fleeter, ¶District Characteristics and Education Costs: Implications of Compensating Wage Differentials on State Aid in California, mimeo (Ohio State University, 1990); and Wayne Wendling, ¶The Cost of Education Index: Measurement of Price Differences of Education Personnel among New York State School Districts, *Journal of Education Finance*, 6 (Spring 1981), pp. 485-504.
29. Monk and Walker (p. 174) argue that a more comprehensive approach "presupposes an ability to reach agreement about the nature and level of outcomes schools are expected to produce." We agree that one must select output measures in order to implement equation 5, but we think that

reasonable procedures can be developed for making this selection. Moreover, the fact that a step may be difficult is a poor excuse for not attempting it, particularly when the conceptual case for it is so strong. Finally, one can estimate cost indices without selecting output measures if one substitutes equation 6 into equation 5. See David Monk and Billy Walker, "The Texas Cost of Education Index: A Broadened Approach," *Journal of Education Finance*, 17 (Fall, 1991), pp. 172-192.

30. There were 695 school districts in New York in 1991. Due to missing observations (including New York City and Yonkers), the sample was limited to 631 observations. The remaining sample appears representative of the major regions in New York State.
31. This measure of expenditure excludes transportation expenses because we do not have any data that would allow us to measure the environmental factors that influence the cost of transporting children to school. In addition, most debt service is excluded from approved operating expenses.
32. A few earlier cost studies, such as Kumar, use pupil-teacher ratios as measures of service quality. We regard this variable as an intermediate output, not a final output, which is not appropriate as a measure of S . Several studies attempt to combine service quality measures and enrollment into a composite output measure. This confuses service outcomes with enrollment, which is an environmental cost factor. See, for example, Downes and Pogue, "Adjusting School Aid Formulas for the Higher Cost of Educating Disadvantaged Students," pp. 89-110; Gyimah-Brempong and Gyapong, "Elasticities of Factor Substitution in the Production of Education," pp. 205-217; Callan

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- and Santerre, "The Production Characteristics of Local Public Education: A Multiple Product and Input Analysis," pp. 468-480; Ramesh Kumar, "Economies of Scale in School Operation: Evidence from Canada," *Applied Economics*, 15 (1983), pp. 323-340; and Emmanuel Jimenez, "The Structure of Educational Costs: Multiproduct Cost Functions for Primary and Secondary Schools in Latin America," *Economics of Education Review*, 5 (1986), pp. 25-39.
33. Byron Brown and Daniel Saks, "The Production and Distribution of Cognitive Skills in Schools," *Journal of Political Economy*, 83 (1975), pp. 571-593.
34. Hanushek, "The Economies of Schooling: Production and Efficiency in Public Schools," p. 1186.
35. Outcome measures were screened out if the adjusted R-squared in the demand model was below 0.1. None of the average achievement test scores available for New York school districts had an R-squared of above 0.06. We followed Brown and Saks and tried including standard deviations from standardized tests as outcome measures. None of the standard deviations had an R-squared in the demand model of above 0.02. The poor performance of average test scores as indicators of voter willingness to pay may explain why these variables were not statistically significant when employed by Downes and Pogue. See Brown and Saks, "The Production and Distribution of Cognitive Skills in Schools," pp. 571-593; and Downes and Pogue, "Adjusting School Aid Formulas for the Higher Cost of Educating Disadvantaged Students," pp. 89-110.

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36. Due to the nature of DEA, it was necessary to convert all outcome measures so that a higher number indicates improved performance. The Regents diploma is awarded to students who pass a relatively difficult set of competency exams in different subject areas. Because not all students are required to take Regency exams, it was not possible to use these test scores directly as outcomes due to sample selectivity problems. Student test scores and drop-out rates are reported in the "Comprehensive Assessment Report," (Albany: New York Department of Education, selected years).
37. A principal component analysis with a varimax rotation was performed on the 18 remaining outcome measures. The eigenvalues of the correlation matrix and the scree plot indicated three distinct factors. The outcomes with high factor scores are most Regents exams and the Regents diploma for factor 1, the dropout rate and some other measures of secondary education for factor 2, and the PEP scores for factor 3. Outcome measures are either based on an average of these measures (PEP scores) or the measure we felt was the best summary measure for the category (Regents diploma and drop-out rate).
38. Teacher salaries are highly related with other professional salaries in New York school districts. The correlation is 0.7 or higher with salaries for principals, assistant principals and superintendents. Salary information on non-professional staff is not available. Salaries and teacher characteristics are collected in the "Personnel Master File" of the "Basic Education Data System" (BEDS) (Albany:

New York Department of Education, selected years). BEDS is a self-reporting survey completed by professional staff in schools. Salaries were adjusted to control for teacher characteristics. To be specific, our salary variable is the residual from a regression of teacher salaries on years of experience, level of education, type of certification, and tenure. A number of districts were missing information on salary levels. We filled in for these missing observations by assuming that a district had the same average adjusted salary level as other districts of the same type (e.g., suburban, rural) in its county.

39. To the best of our knowledge, only one previous study, Downes and Pogue, recognizes that teacher wages are endogenous. However, their study fails to eliminate endogeneity bias because one of the instruments in their simultaneous equations procedure is an index of teacher experience, which also is endogenous. See Downes and Pogue, "Adjusting School Aid Formulas for the Higher Cost of Educating Disadvantaged Students," pp. 89-110.
40. Economies to pupil scale need to be distinguished from economies to quality scale and economies of scope. See Duncombe and Yinger, "An Analysis of Returns to Scale in Public Production, with an Application to Fire Protection," pp. 49-72.
41. Because we use a double-log functional form, we actually include the log of enrollment and the square of the log of enrollment. Either enrollment or average daily attendance, ADA, could be used as the measure of the number of pupils. An argument can be made for each as the most directly

related to costs. We selected enrollment since school districts are likely to budget resources for close to full attendance. However, the correlation between enrollment and ADA is close to 1.0 in New York and there was little change in the cost indices when ADA was used. See Monk, *Educational Finance: An Economic Approach*.

42. See, for example, Ratcliffe, Riddle, and Yinger, "The Fiscal Condition of School Districts in Nebraska: Is Small Beautiful?" pp. 81-99; and Downes and Pogue, "Adjusting School Aid Formulas for the Higher Cost of Educating Disadvantaged Students," pp. 89-110.
43. See R. Gary Bridge, Charles Judd, and Peter Moock, *The Determinants of Educational Outcomes: The Impact of Families, Peers, Teachers, and Schools* (Ballinger Publishing Company, 1979), and Hanushek, "The Economics of Schooling: Production and Efficiency in Public Schools," pp. 1141-1177. These variables and others are discussed in the reviews mentioned earlier. One example of a production study that uses all three of these environmental factors is Ronald Ferguson, "Paying for Public Education: New Evidence on How and Why Money Matters," *Harvard Journal on Legislation*, 28 (1991), pp. 465-498.
44. The source of most of these variables is the *1990 Census* as reported in the "School District Data Book" (Washington, DC: U.S. Bureau of the Census and the National Center for Education Statistics, 1994). The remaining variables come from the New York Department of Education's Basic Education Data System.

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45. See, particularly, Ladd and Yinger, *America's Ailing Cities: Fiscal Health and the Design of Urban Policy*; Daniel Rubinfeld, "The Economics of the Local Public Sector," in A. Auerback and M. Feldstein, eds., *Handbook of Public Economics*, Vol. 2 (New York: Elsevier Science Publishers, 1985), pp. 571-645; and Robert Inman, "The Fiscal Performance of Local Governments: An Interpretative Review," in P. Mieszkowski and M. Straszheim, eds., *Current Issues in Urban Economics* (Baltimore: The Johns Hopkins University Press, 1979), pp. 270-321.
46. While on paper the operating aid formula used in New York is similar in design to a matching percent equalizing grant, in actual practice it is closer to a lump-sum foundation grant. Since aid is lump-sum and is distributed based on a measure of fiscal capacity, it is likely to be exogenous to local district spending decisions. See Miner for a good discussion of school aid formulas in New York State. See Jerry Miner, "A Decade of New York State Aid to Local Schools," Metropolitan Studies Program occasional Paper No. 141, Center for Policy Research, The Maxwell School (Syracuse, NY: Syracuse University, 1991).
47. Borrowing from Ladd and Yinger, and Duncombe, the tax share is represented as; $V_m/V = (V_m/V_1)(1-e)$, where V_m and V_1 are the median and average local residential property values and e is the percent of property taxes borne by non-residents. We construct the export ratio, e , using information on the distribution of property values by type (from the New York State

Department of Equalization and Assessment) and estimates of property tax exporting by type of property from Ladd and Yinger. Similar results are obtained using a set of property composition variables instead of the export ratio. See Ladd and Yinger, *America's Ailing Cities: Fiscal Health and the Design of Urban Policy*; William Duncombe, *Demand for Local Public Services Revisited: The Case of Fire Protection*, *Public Finance Quarterly*, 19 (1991), pp. 412-436;

48. One could argue that the percentage of adults with a college degree is an environmental cost variable; more educated parents do more to reinforce the lessons their children learn in school. When this variable is treated as an environmental cost factor, however, it has the wrong sign (positive) so we cannot reject the hypothesis that it has no impact on costs. A similar procedure ruled out children per household and median income as cost variables.
49. See, for example, W.A. Niskanen, *Bureaucracy and Representative Government* (Aldine-Atherton, 1971); W.A. Niskanen, *Bureaucrats and Politicians*, *The Journal of Law and Economics*, 18 (1975), pp. 617-643.; Paul G. Wyckoff, *A Bureaucratic Theory of Flypaper Effects*, *Journal of Urban Economics*, 23 (1988), pp. 115-129; and Paul G. Wyckoff, *The Simple Analytics of Slack-Maximizing Bureaucracy*, *Public Choice*, 67 (1990), pp. 35-67.
50. See Duncombe, Miner and Ruggiero for a more complete discussion of factors associated with cost inefficiency. Contrary to expectation, they found a negative relationship between the relative number of private school students (or schools) and the level of cost efficiency. William Duncombe,

Jerry Miner, and John Ruggiero, "Empirical Evaluation of Bureaucratic Models of Inefficiency," *Public Choice* (1995), forthcoming.

51. The Cobb-Douglas cost function restricts the elasticity of substitution between all factor inputs to be one and assumes homotheticity between costs and outputs. Since we only include one factor price in the cost model, the factor substitution restriction is not a serious limitation.
52. Another approach is to estimate a translog or flexible functional form. This approach is taken by Jimenez, Callan and Santerre, and Gyimah-Brempong and Gyapong. Duncombe and Yinger have used this approach in the production context but in this case we believe that it would add complexity without significant insight. See Jimenez, "The Structure of Educational Costs: Multiproduct Cost Functions for Primary and Secondary Schools in Latin America," pp. 25-39; Callan and Santerre, "The Production Characteristics of Local Public Education: A Multiple Production and Input Analysis," pp. 468-480; Gyimah-Brempong and Gyapong, "Elasticities of Factor Substitution in the Production of Education," pp. 205-217; and Duncombe and Yinger, "An Analysis of Returns to Scale in Public Production, With an Application to Fire Protection," pp. 49-72.
53. These results suggest that if consolidation of small districts is not possible, it is appropriate to control for the cost effects of scale in an education cost index. See Duncombe, Miner, and Ruggiero for an analysis of the benefits of school district consolidation in New York State. They

found that the number of districts that might benefit from consolidation in New York and the potential cost savings from consolidation were quite small. William Duncombe, Jerry Miner, and John Ruggiero, *Potential Cost Savings from School District Consolidation: A Case Study of New York*, *Economics of Education Review* (1995), forthcoming.

54. Categories of handicapped students are organized by the level of special services they receive. Categories include students requiring special services 60 percent or more of the day or using private schools for services, students requiring special services at least 20 percent of the day, student requiring consultant teacher services, and students who use two periods a week in special services. Several New York State education aid formulas use total weighted pupil units, TWPU, which assigns different pupil weightings to these handicapping categories.
55. Since the price of labor is treated as endogenous in the cost model, a predicted wage is used to construct the cost index. The predicted wage is based on the predicted value of a first-stage regression between the price of labor and all exogenous and instrumental variables used in the cost model.
56. These cost index calculations are similar to the ones used by Downs and Pogue, *Adjusting School Aid Formulas for the Higher Cost of Educating Disadvantaged Students*, pp. 89-110, although, as explained earlier, our cost model differs from theirs in several respects.

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57. This approach has been applied to education costs by Ratcliffe, Riddle, and Yinger, *The Fiscal Condition of School Districts in Nebraska: Is Small Beautiful?*, and Downes and Pogue, *Adjusting School Aid Formulas for the Higher cost of Educating Disadvantaged Students*, pp. 89-110.
 58. To construct a teacher salary cost index, we regressed actual teacher salaries on factors associated with differences in teacher quality (experience, certification, level of education, and tenure), demand for educational services, county population (as a proxy for private wages) and student and family background characteristics. All factors were held at the state mean except the county population and student and family background characteristics.
 59. We estimated an indirect cost model with efficiency, but none of the variables in the model were statistically significant and we decided not to present the results.
 60. This result contradicts the finding in Downes and Pogue, whose direct and indirect approaches yield cost indices that are highly correlated. This difference may reflect our inability to incorporate efficiency into our indirect approach. Downes and Pogue do not have to deal with this issue because, as noted earlier, they account for efficiency by differencing. However, the Downes and Pogue cost index is based on fewer cost factors, so their direct and indirect cost indices might differ if more factors were included. See Downes and Pogue, *Adjusting School Aid Formulas for the Higher Cost of Educating Disadvantaged Students*, pp. 89-110.

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61. Part of the reason that the teacher salary index does not demonstrate a U-shape is because we do not include enrollment variables in this cost model. While some studies in the past have included enrollment, no scholar has provided a convincing reason why teacher wages should be directly related to variation in enrollment. Interestingly, all these studies find an inverted U-shaped function between enrollment size and salaries. See Chambers, "Educational Cost Differentials and the Allocation of State Aid for Elementary and Secondary Education," pp. 459-481; Wendling, "The Cost of Education Index: Measurement of Price Differences of Education Personnel among New York State School Districts," pp. 485-504; and Fleeter, "District Characteristics and Education Costs: Implications of Compensating Wage Differentials on State Aid in California."
62. The concepts used in DEA were conceptualized by Farrel and developed by Charnes, Cooper and Rhodes, and Färe and Lovell to analyze multiple output production correspondences. See M.J. Farrell, "The Measurement of Productive Efficiency," *Journal of the Royal Statistical Society, Series A, General*, 120 (1957), pp. 253-281; Rolf Färe and Knox Lovell, "Measuring the Technical Efficiency of Production," *Journal of Economic Theory* (1978), pp. 150-162; and A. Charnes, W.W. Cooper, and E. Rhodes, "Measuring the Efficiency of Decision Making Units," *European Journal of Operational Research*, 2 (1978), pp. 429-444.
63. For a further discussion of strengths and weaknesses of DEA, see L. Seiford and R. Thrall, "Recent Developments in DEA: The Mathematical Programming Approach to Frontier Analysis," *Journal*

of Econometrics, 46 (1990), pp. 7-38.

64. See Grosskopf and Yaisawarng for one of the first applications of DEA to cost frontiers. Grosskopf and Yaisawarng limit their sample so that all producing units face the same cost environment. See Grosskopf and Yaisawarng, "Economies of Scope in the Provision of Local Public Services," pp. 61-74.
65. See, for example, A. Bessent and E.W. Bessent, "Determining the Comparative Efficiency of Schools through Data Envelopment Analysis," *Educational Administration Quarterly*, 16(1980), pp. 57-75; and Rolf Färe, Shawna Grosskopf, and William Weber, "Measuring School District Performance," *Public Finance Quarterly*, 17 (1989), pp. 409-428.
66. See Ruggiero, "Are Costs Minimized in the Public Sector? A Nonparametric Analysis of the Provision of Educational Services," and John Ruggiero, "On the Measurement of Technical Efficiency in the Public Sector," *European Journal of Operational Research*, (1995), forthcoming.

Table 1. Descriptive Statistics for Cost Model and Instruments

(New York school districts in 1991, n=631)

| Variable | Mean | Standard deviation | Minimum | Maximum |
|--|--------|--------------------|---------|----------|
| Cost Model: | | | | |
| Log of per pupil expenditures | 8.66 | 0.29 | 8.06 | 10.14 |
| PEP scores (average percent of students above SRP) | 94.24 | 3.79 | 64.50 | 100.00 |
| Percent receiving Regents diploma | 40.44 | 13.07 | 0.00 | 75.38 |
| Percent non-dropouts | 97.59 | 1.84 | 88.10 | 100.00 |
| Log of teacher salaries | 10.11 | 0.12 | 9.56 | 10.46 |
| Log of enrollment | 7.37 | 0.89 | 4.36 | 10.75 |
| Percent of children in poverty | 11.57 | 7.45 | 0.26 | 38.04 |
| Percent female-headed households | 8.79 | 2.71 | 2.46 | 34.68 |
| Percent handicapped students | 10.64 | 3.37 | 1.63 | 30.68 |
| Percent severely handicapped students | 4.49 | 2.12 | 0.00 | 14.57 |
| Persons with limited English proficiency (percent) | 0.99 | 1.27 | 0.00 | 11.96 |
| High school students (percent) | 28.97 | 3.71 | 20.09 | 63.10 |
| Efficiency index (percent)^a | 66.46 | 15.76 | 19.49 | 100.00 |
| Instruments: | | | | |
| Log of median family income | 10.55 | 0.31 | 9.96 | 11.63 |
| Log of operating aid | 7.53 | 0.50 | 6.06 | 8.18 |
| Log of tax share | 0.05 | 0.22 | -0.52 | 0.70 |
| Percent owner-occupied housing | 75.36 | 10.16 | 36.50 | 95.38 |
| Percent of households with children | 33.41 | 5.29 | 19.14 | 52.80 |
| Percent of adults with college education | 19.08 | 11.05 | 4.05 | 69.66 |
| City district (1=yes) | 0.09 | 0.29 | 0.00 | 1.00 |
| 1990 county population (thousands) | 388.94 | 457.03 | 5.28 | 1,321.86 |

Source: New York State Department of Education, *Comprehensive Assessment Report, Basic Education Data System and Fiscal Profile*, and National Center for Education Statistics, *School District Data Book*.

^aEfficient districts have an index of 100. Based on DEA estimates for the three outcome variables listed and per pupil expenditures.

Table 2. Education Cost Model Results—New York School Districts (1991)(Regression coefficients, n=631)^a

| Variables | Three outcomes | | Two outcomes | |
|--|--------------------|--------------------|---------------------|--------------------|
| | (Model 1) | (Model 2) | (Model 3) | (Model 4) |
| Intercept | -7.7095 (-2.67) | -8.0172 (-2.68) | 1.5291 (0.92) | -1.8377 (-0.56) |
| PEP scores (average percent above SRP) | 2.3472 (1.49) | 2.3986 (1.52) | 2.1877 (3.24) | 2.9261 (2.11) |
| Percent non-dropouts | 7.1626 (2.68) | 6.4159 (2.35) | 5.2284 (4.00) | |
| Percent receiving Regents diploma | 1.2432 (2.73) | 1.3156 (2.85) | | 1.5275 (3.78) |
| Efficiency index (percent) | -0.9930 (-4.49) | -0.9337 (-4.02) | -1.5660 (-13.84) | -1.1436 (-5.09) |
| Log of teacher salaries | 0.8913 (2.16) | 0.9936 (2.38) | 0.1530 (0.92) | 0.9657 (2.63) |
| Log of enrollment | -0.5331 (-3.73) | -0.5552 (3.75) | -0.2503 (-3.76) | -0.5397 (-4.02) |
| Square of log of enrollment | 0.0329 (3.81) | 0.0338 (3.78) | 0.0163 (3.68) | 0.0309 (3.70) |
| Percent of children in poverty | 0.8306 (3.99) | 0.7903 (3.76) | 0.4812 (4.53) | 0.5036 (2.69) |
| Percent female-headed households | 2.1166 (4.26) | 1.9823 (3.95) | 0.6033 (2.36) | 1.7162 (3.82) |
| Percent handicapped students | 0.3903 (1.11) | | | |
| Percent severely handicapped students ^b | | 0.9656 (1.66) | 0.5295 (1.88) | 0.4460 (0.93) |
| Persons with limited English proficiency (percent) | 2.5236 (2.11) | 2.5844 (2.10) | 1.3943 (2.26) | 3.0664 (2.84) |
| High school students (percent) | 0.2945 (1.10) | 0.3438 (1.26) | 0.4451 (3.34) | 0.3756 (1.51) |
| SSE | 24.82 | 26.05 | 6.50 | 22.08 |
| Adjusted R-square | 0.51 | 0.48 | 0.87 | 0.56 |

^aCost model estimated with linear 2SLS regression using instruments reported in Table 1. The cost model is based on a modified Cobb-Douglas production function with the square of enrollment. The dependent variable is the log of per pupil approved operating expenditures. T-statistics are in parentheses.

^bStudents are in special class instruction or special programs for at least 60 percent of school day.

Table 3. Correlations between Education Cost Indices for New York State School Districts in 1991^a

(n=631)

| Socio-economic characteristics | Direct cost indices | | | Indirect cost index (No efficiency index) | Cost index based on weighted pupils | Teacher salary cost index |
|---|-----------------------|---------------------|----------------------|--|-------------------------------------|---------------------------|
| | Endogenous efficiency | No efficiency index | Exogenous efficiency | | | |
| Standard deviation | 16.93 | 26.11 | 10.85 | 15.33 | 8.88 | 12.75 |
| Maximum | 239.62 | 356.11 | 191.84 | 253.94 | 264.00 | 143.55 |
| 75th percentile | 105.26 | 109.53 | 103.21 | 105.74 | 102.94 | 111.61 |
| 25th percentile | 89.56 | 83.89 | 93.77 | 90.50 | 96.96 | 89.81 |
| Minimum | 77.50 | 70.70 | 83.83 | 77.05 | 44.68 | 68.37 |
| Correlations: | | | | | | |
| Direct Cost Indices: | | | | | | |
| Endogenous efficiency index | 1.00 | | | | | |
| No efficiency index | 0.94 | 1.00 | | | | |
| Exogenous efficiency index | 0.84 | 0.74 | 1.00 | | | |
| Indirect cost index (no efficiency index) | 0.63 | 0.55 | 0.39 | 1.00 | | |
| Cost index based on weighted pupils | 0.14 | 0.15 | 0.13 | 0.08 | 1.00 | |
| Teacher salary cost index | 0.47 | 0.57 | 0.32 | -0.08 | 0.06 | 1.00 |

^aIndex for first three columns is based on 3-factor cost model (model 2 in Table 2) with the state average equal to 100. Index for fourth column is a reduced form model where the demand instruments—income, taxshare, and households with children—are substituted into the cost model for outcome measures. The index in the fifth column is based on a ratio of weighted pupils over total enrollment; extra weight is given to secondary, handicapped and special needs pupils. The index in the last column is based on the relationship between teacher salaries and family and student characteristics.

Table 4. Comparison of Education Cost Indices^a for New York State School Districts in 1991

(n=631)

| Socio-economic characteristics | Number of districts | Direct cost indices | | | Indirect cost index (no efficiency index) | Cost index based on weighted pupils | Teacher salary cost index |
|--------------------------------|---------------------|-----------------------|---------------------|----------------------|---|-------------------------------------|---------------------------|
| | | Endogenous efficiency | No efficiency index | Exogenous efficiency | | | |
| Region type: | | | | | | | |
| Downstate small cities | 7 | 130.0 | 142.0 | 112.3 | 105.9 | 102.6 | 124.3 |
| Downstate suburbs | 130 | 111.4 | 125.0 | 102.0 | 100.9 | 101.6 | 117.5 |
| Upstate large cities | 3 | 179.2 | 190.2 | 172.6 | 132.8 | 100.3 | 124.5 |
| Upstate rural | 212 | 98.4 | 93.5 | 99.1 | 107.1 | 99.9 | 90.1 |
| Upstate small cities | 48 | 105.5 | 101.2 | 111.8 | 95.2 | 100.6 | 99.6 |
| Upstate suburbs | 231 | 91.9 | 89.2 | 95.9 | 93.4 | 99.0 | 98.3 |
| Pupil size class: | | | | | | | |
| Under 100 pupils | 1 | 156.5 | 166.1 | 109.8 | 253.9 | 120.3 | 86.0 |
| 100-500 pupils | 61 | 108.8 | 110.8 | 100.6 | 128.4 | 101.1 | 89.9 |
| 500-1,000 pupils | 113 | 101.0 | 100.1 | 98.7 | 106.8 | 99.4 | 94.7 |
| 1,000-1,500 pupils | 131 | 94.6 | 92.2 | 97.0 | 96.6 | 98.5 | 96.1 |
| 1,500-3,000 pupils | 182 | 96.7 | 95.9 | 98.5 | 93.2 | 100.7 | 101.9 |
| 3,000-5,000 pupils | 80 | 97.2 | 99.6 | 99.8 | 90.2 | 99.6 | 107.3 |
| 5,000-10,000 pupils | 54 | 108.8 | 111.4 | 108.8 | 94.9 | 101.1 | 111.8 |
| Over 10,000 pupils | 9 | 139.4 | 149.0 | 133.3 | 110.9 | 100.9 | 119.3 |

Table 4. Continued

| Socio-economic characteristics | Number of districts | Direct cost indices | | | Indirect cost index (no efficiency index) | Cost index based on weighted pupils | Teacher salary cost index |
|--------------------------------------|---------------------|-----------------------|---------------------|----------------------|---|-------------------------------------|---------------------------|
| | | Endogenous efficiency | No efficiency index | Exogenous efficiency | | | |
| Income class (percentile): | | | | | | | |
| Under 10th | 62 | 106.3 | 118.5 | 97.7 | 100.8 | 100.5 | 115.5 |
| 10th to 25th | 95 | 99.8 | 105.4 | 97.3 | 100.8 | 100.2 | 109.4 |
| 25th to 50th | 157 | 98.5 | 98.7 | 100.2 | 103.2 | 99.4 | 101.5 |
| 50th to 75th | 159 | 98.7 | 95.7 | 100.8 | 111.4 | 99.5 | 95.0 |
| 75th to 90th | 94 | 99.0 | 94.1 | 101.3 | 76.8 | 100.6 | 92.1 |
| Over 90th | 64 | 102.6 | 96.6 | 101.9 | 74.8 | 101.0 | 91.2 |
| Property values (percentile): | | | | | | | |
| Under 10th | 63 | 109.7 | 122.9 | 98.7 | 111.7 | 100.1 | 109.1 |
| 10th to 25th | 94 | 105.2 | 112.0 | 99.1 | 100.4 | 100.1 | 109.1 |
| 25th to 50th | 158 | 99.8 | 100.5 | 100.0 | 97.6 | 99.5 | 101.9 |
| 50th to 75th | 158 | 94.1 | 89.6 | 97.9 | 97.6 | 100.0 | 95.2 |
| 75th to 90th | 95 | 96.7 | 91.0 | 100.7 | 98.7 | 99.7 | 94.5 |
| Over 90th | 63 | 103.0 | 97.5 | 106.8 | 101.7 | 101.3 | 92.9 |

^aIndex for first three columns is based on 3-factor cost model (model 2 in Table 2) with the state average equal to 100. Index for fourth column is a reduced form model where the demand instruments—income, tax share and households with children—are substituted into the cost model for outcome measures. The index in the fifth column is based on a ratio of weighted pupils over total enrollment; extra weight is given to secondary, handicapped and special needs pupils. The index in the last column is based on the relationship between teacher salaries and family and student characteristics. Income is based on estimated per capita adjusted gross income in 1991 and property values are per capita market value for all property in 1990.

Table 5. Impact of Input and Environmental Variables on Education Costs, Districts with Highest and Lowest Costs^a

| Cost variables | 10 percent of districts with highest costs | | 10 percent of districts with lowest costs | |
|--|--|--|---|--|
| | Index relative to state average (=100) | Percent of cost difference due to variable | Index relative to state average (=100) | Percent of cost difference due to variable |
| Total cost index | 152.7 | | 81.1 | |
| (Per pupil difference from average district) | \$3,046 | | -\$1,091 | |
| Teacher salaries | 111.2 | 30.67% | 96.0 | -20.03% |
| Log of enrollment | 198.3 | -13.20% | 128.0 | -18.85% |
| Percent of children in poverty | 161.6 | 13.98% | 49.4 | -22.86% |
| Percent female-headed households | 170.9 | 36.16% | 75.3 | -21.24% |
| Percent severely handicapped students ^b | 177.7 | 9.43% | 54.3 | -9.90% |
| Limited English proficiency (percent) | 407.6 | 22.47% | 39.3 | -7.77% |
| High school students (percent) | 94.9 | -1.38% | 101.3 | 0.64% |

^aCost index based on 3-factor cost model (model 2 in Table 2). Indices for costs and environmental variables relative to state average are based on average values for 10 percent of districts with the highest and lowest per pupil costs. The percentage cost difference due to cost variable i , say Pc_i , is based on three cost indices. Index A is the total cost index for the high-(or low-) cost districts. Index B is a cost index with all variables set at the state average. Index C is a cost index with cost variable i set at the average for the high-(or low-) cost districts and all other variables set at the state average. Then, $Pc_i = (C - B) / (A - B)$.

^bStudents are in special class instruction or special programs for at least 60 percent of school day.

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